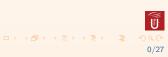
Realized Networks

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Introduction



High Frequency Data Based Volatility Estimation

- Over the last decade, the availability of intra-daily high frequency trade, quote and order book data has boosted research on the construction of efficient ex-post measure of daily return variability
- These estimator are typically called realized volatility estimators
- Extensive literature on the topic: Andersen, Bollerslev, Diebold and Labys (2003); Ait-Sahalia, Mykland and Zhang (2005); Bandi and Russell (2006); Barndorff-Nielsen, Hansen, Lunde, and Shephard (2009); and many others

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- Besides numerical challenges, realized covariance estimation suffers from two challenges which are inherently linked to covariance estimation for large number of assets:

Precise estimation of the covariance (cf Ledoit & Wolf, 2004



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 Interpretation of the dependence structure of the assets



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 - 1 Precise estimation of the covariance (cf Ledoit & Wolf, 2004)
 - **2** Interpretation of the dependence structure of the assets

- We propose a LASSO-based regularization procedure for realized covariance estimation.
 - It shrinks the off diagonal elements of the inverse realized covariance to zero
 - Regularized estimator can be interpreted as a partial correlation network
 We call our estimator the Realized Network
- 2 We analyse the large sample properties of the estimator.
 - \blacksquare We employ standard covariance estimation framework which allows for
 - market-microstructure noise and asynchronous trading
 - Focus on Two Scales Realized Covariance estimator and Multivariate Realized Kernel Establish conditions of consistent covariance estimation and network selection
- Advantages of the methodology are illustrated by means of a simulation study and an empirical illustration

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Hautsch, Kyj and Oomen (2012); Corsi, Peluso, and Audrino (2015); Malec, Hautsch, Kyj (2015); Wang and Zhou (2010); Tao, Wang and Zhou (2013);

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Framework Assumptions Realized Covariance Estimator Realized Network Estimator

Framework



Framework Assumptions Realized Covariance Estimator Realized Network Estimator Covariance of the Efficient Price

• y(t) denotes the efficient log-price of *n* assets.

• y(t) is a Brownian martingale

$$y(t)=\int_0^t \Theta(u)dB(u) ,$$

where B(u) Brownian motion and $\Theta(u)$ is the spot covolatility

Integrated Covariance: the covariance matrix of daily return y(1)

$$\operatorname{Var}(y) = \int_0^1 \Sigma(t) dt = \Sigma^*$$

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where $\Sigma(t) = \Theta(t)\Theta(t)'$



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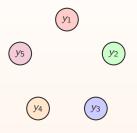
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- The network associated with the system is an undirected graph



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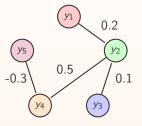
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Partial Correlation measures (cross-sect.) linear conditional dependence between y_i t and y_j t given on all other variables:

$$\rho^{ij} = \mathsf{Cor}(y_{i\,t}, y_{j\,t} | \{y_{k\,t} : k \neq i, j\}).$$

 Partial Correlation is related to Linear Regression: For instance, consider the model

 $y_{1t} = c + \beta_{12}y_{2t} + \beta_{13}y_{3t} + \beta_{14}y_{4t} + \beta_{15}y_{5t} + u_{1t}$

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Network is entirely characterized by the integrated concentration matrix K^{*} = (Σ^{*})⁻¹ = (k^{*}_{ij}):

$$\rho^{ij} = \frac{-k_{ij}^{\star}}{\sqrt{k_{ii}^{\star}k_{jj}^{\star}}}$$

In particular, the nonzero entries of \mathbf{K}^* correspond to the linkages of the network.

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- In finance, it is customary to assume that returns have a factor structure. In practice, it is more interesting to analyse the partial correlation structure of assets conditional on the factors
- Thus, in this work we define the network on the basis of the idiosyncratic integrated covariance, that is the covariance of the residuals obtained by projecting the factors onto the assets

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It is customary to assume that the econometrician does not observe the efficient price y but a contaminated version x defined as

$$x_i(t_{i\,k}) = y_i(t_{i\,k}) + u_i(t_{i\,k})$$

where

t_{ik} (asset specific) timestamp of a trade/midquote
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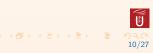


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Objectives:

- **1** Estimate the integrated covariance
- Detect the nonzero linkages of the network, which is equivalent to detecting the nonzero entries of the integrated concentration matrix.
- We are going to tackle both objectives simulatenously by introducing a sparse integrated concentration matrix estimator. More specifically:

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Framework Assumptions Realized Covariance Estimator Realized Network Estimator Two Scales Realized Covariance Estimator

- Our estimation strategy consist of regularizing a consistent RC estimator. Many estimators are available in the setting we are working on
- In this presentation we focus on the Two Scales Realized Covariance estimator ∑ based on Pairwise–Refresh Time (TSRC). The estimator deals with both market microstructure noise and asynchronicity.
- Other sufficiently regular estimators can be used. Theory does not depend non specific form of the estimator. In particular, the paper also develops the theory for the Multivariate Realized Kernel.



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Framework Assumptions Realized Covariance Estimator Realized Network Estimator Two–Scale Realized Covariance (TSRC)

Let (x^r_{ik}, x^r_{jk}) denote the "pairwise refresh time" adjusted observed prices for stock i and j

• The TSRC estimator is denoted by $\overline{\Sigma}_{TS} = (\overline{\sigma}_{TS,ij})$,

$$\overline{\sigma}_{\mathsf{TS},ij} = \frac{1}{K} \sum_{k=K+1}^{m} \left(x_{i\,k}^{r} - x_{i\,k-K}^{r} \right) \left(x_{j\,k}^{r} - x_{j\,k-K}^{r} \right) - \frac{m_{K}}{m_{J}} \frac{1}{J} \sum_{k=J+1}^{m} \left(x_{i\,k}^{r} - x_{i\,k-J}^{r} \right) \left(x_{j\,k}^{r} - x_{j\,k-J}^{r} \right)$$

where $m_K = \frac{m-K+1}{K}$ and $m_J = \frac{m-J+1}{J}$.

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The Realized Network Estimator is defined as

$$\widehat{\mathbf{K}} = \arg\min_{\mathbf{K}\in\mathcal{S}^n} \left\{ \operatorname{tr}(\overline{\Sigma}\mathbf{K}) - \log\det(\mathbf{K}) + \lambda \sum_{i\neq j} |k_{ij}| \right\}$$

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Theory

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Key ingredient to establish results is a concentration inequality of the realized volatility estimator.

Define *M* as the minimum sample size used to compute a realized covariance entry across all pairs of stocks.

Then we have

 $\mathbb{P}\left(\left|\overline{\sigma}_{ij} - \sigma_{ij}^{\star}\right| > x\right) \le a_1 M^{\alpha} \exp\left\{-a_2 \left(M^{\beta} x\right)^{\gamma}\right\}.$

for some positive exponents $lpha,eta,\gamma$



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Consistent Estimation

Theorem: Consistent Concentration Estimation

Let
$$\lambda = \frac{8}{\alpha} M^{-\beta} \left(\frac{\log(a_1 n^{\tau})}{a_2} \right)^{\frac{1}{\gamma}}$$
 for some $\tau > 2$.

Let

$$M > C\left(\log\left(a_2(a_1^{\frac{1}{\beta\gamma}}C_0(d)^{\frac{1}{\beta}})n^{\tau}\right)\right)^{\frac{1}{\beta\gamma}}C_0(d)^{\frac{1}{\beta}},$$

where C_0 is a function of the max vertex degree d

Then, for n sufficiently large

$$\Pr\left(||\widehat{\mathbf{K}} - \mathbf{K}^{\star}||_{\infty} \leq 2C_{\mathbf{\Gamma}^{\star}} \left(1 + \frac{8}{\alpha}\right) M^{-\beta} \left[\frac{\log\left(a_{1}M^{\alpha}n^{\tau}\right)}{a_{2}}\right]^{\frac{1}{a_{0}}}\right) \geq 1 - \frac{1}{n^{\tau-2}}$$

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where C_{Γ^*} is a constant that depends on Σ .

Theorem: Consistent Network Selection

Let
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where C_1 is a function of the max vertex degree d. Then, for n sufficiently large

$$\mathrm{P}\left(\mathrm{sign}(\widehat{k}_{ij}) = \mathrm{sign}(k_{ij}^{\star}), \forall i, j \in \{1, \dots, n\}\right) \geq 1 - \frac{1}{n^{\tau-2}}.$$

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- It is useful to analyse how the expression simplify depending on the degree of sparsity of the networks
 - If the max degree d is zero (disconnected graph), then the sample size M has to be at least O((log n)^{1/βγ})
 - 2 If the max degree *d* is O(n) (fully connected graph), then the sample size *M* has to be at least $O((\log n)^{\frac{1}{\beta\gamma}}n^{\frac{1}{\beta}})$
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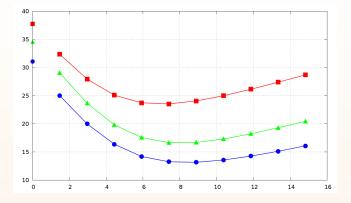


Simulation Study

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Simulation Study Peek: MSE vs Network Shrinkage



- Simulation study to analyse finite sample properties of the procedure.
- In particular, study shows that if partial correlation structure is sparse, realized network estimator substantially enhances precision

Empirical Application



- We consider a panel of 96 NYSE Bluechips (\approx constituents of the S&P 100)
- We estimate realized covariance for each day of 2009
- Realized covariance is estimated using the Realized Network estimator based on TSRC. (tuning parameter λ chosen via the BIC)
- Estimators are computed using trade prices from the NYSE-TAQ Standard procedures are applied to clean and filter the data
- We focus on the idiosyncratic covariance matrix We analyse interdependence conditional on the market factor.



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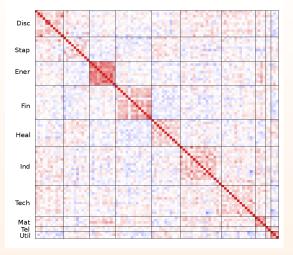


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Realized Network Estimates

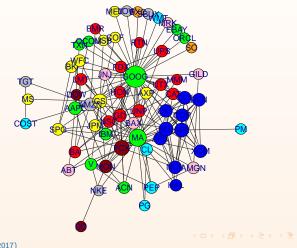
Realized Correlation Heatmap on 2009-07-02



Brownlees, Nualart & Sun (2017)

Realized Network Estimates

Realized Network on 2009-07-02

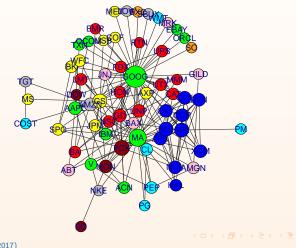


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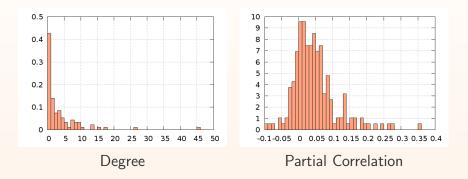
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Empirical Application Degree and Partial Correlation Distribution





GMV portfolio prediction exercise:

- Construct the GMV portfolio weights using the MRK
 - Competitors: Unconstrained, Constrained, Shrinkage and Realized Network
- 2 Use the weights to construct daily GMV portfolio for the following day.
- 3 Compute the variance of the daily portfolios over the full year

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- More precise covariance estimators deliver GMV portfolio weights that generate smaller out-of-sample portfolio variances (cf Engle and Colacito, 2006)

Predictive Analysis: GMV Comparison

	No Regular.	Diagonal	Network	Factor	Shrinkage	Block-Factor
RC	13.40	13.74	12.09	12.11	12.03	13.24
		-0.70	1.41	2.46**	2.65***	0.26
TSRC	13.05	13.52	11.96	12.34	12.16	12.29
		-0.74	1.88*	1.34	1.73*	0.88
MRK	12.95	13.89	11.65	13.03	10.46	11.15
		-1.82^{*}	2.17**	-0.12	4.19***	2.30**

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- Highlights:
 - The procedure delivers more precise estimates of the covariance when the partial correlation structure of the assets is sparse.
 Regularized estimator can be represented as a network.
- Empirical application shows that regularization significantly improves the estimator. In a GMV portfolio forecasting exercise, substantial gains in prediction accuracy



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Questions?

Thanks!

